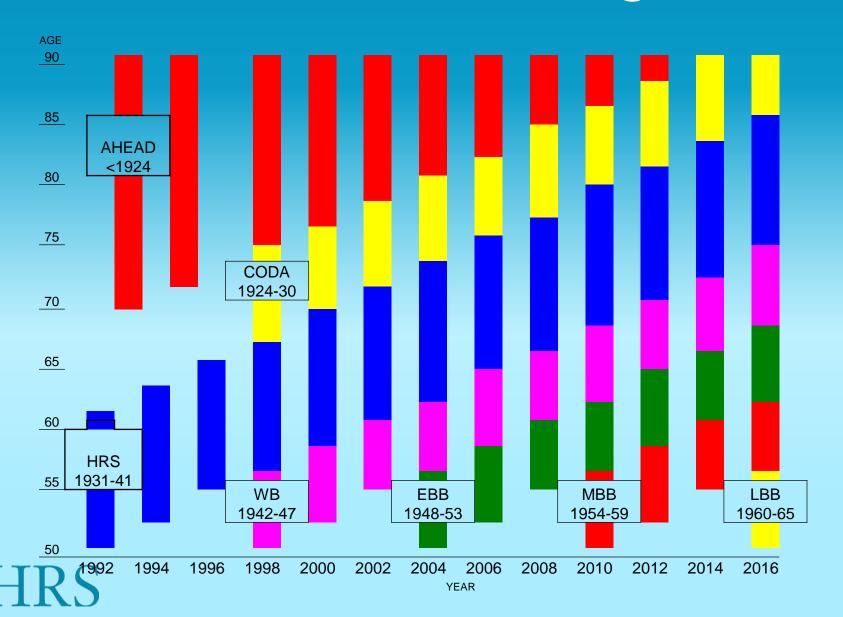
Panel 5: Cohort Changes

Cohort Changes in Social Security and Pension Wealth

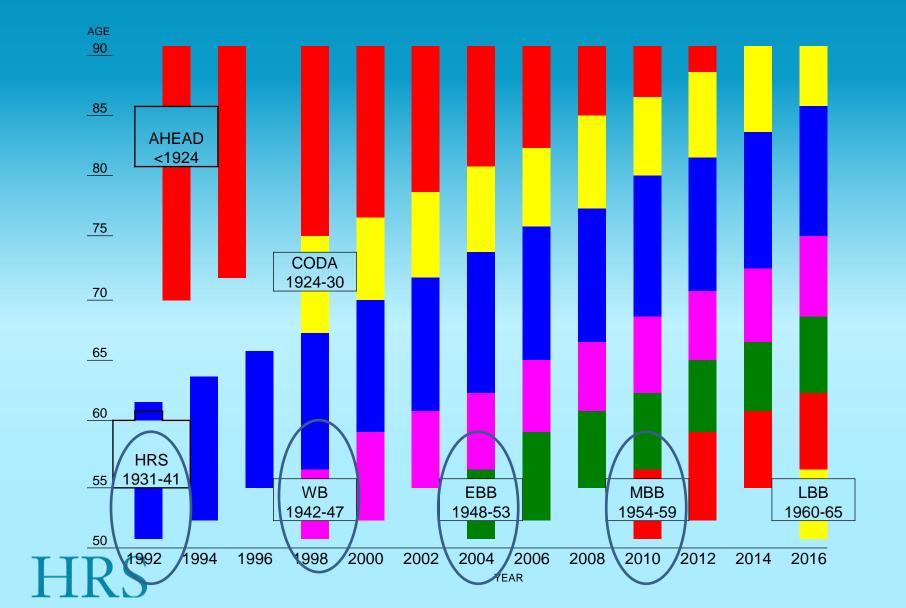
Chichun Fang, Charles Brown, and David Weir HRS, Institute for Social Research, University of Michigan



HRS Cohort Design



HRS Cohort Design



Crucial question for an aging America:

As we move from the small cohorts born before 1946 to the large ones of the baby boom, how is their health and preparation for retirement changing?



For middle-aged whites, US mortality headed in the wrong direction

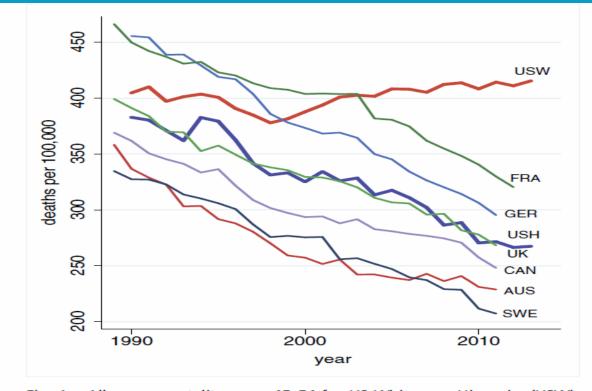
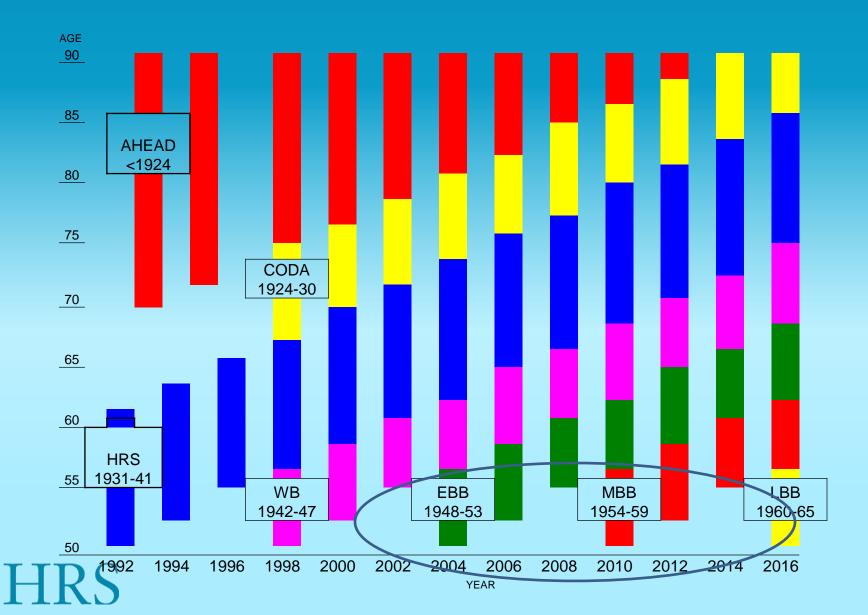


Fig. 1. All-cause mortality, ages 45–54 for US White non-Hispanics (USW), US Hispanics (USH), and six comparison countries: France (FRA), Germany (GER), the United Kingdom (UK), Canada (CAN), Australia (AUS), and Sweden (SWF)

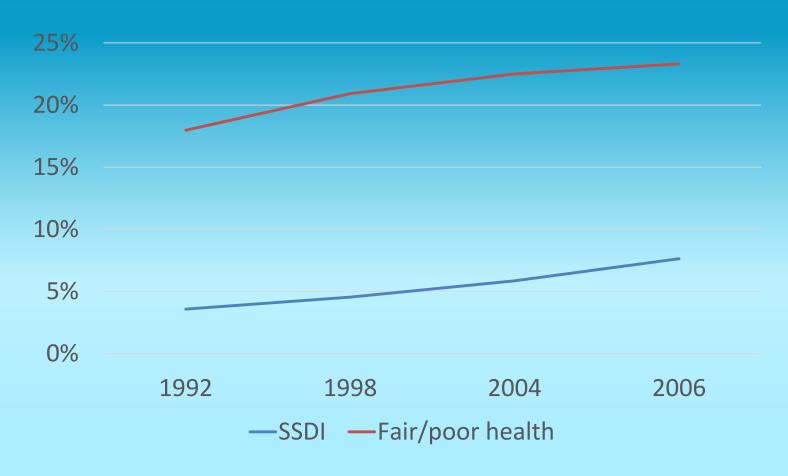


Source: Anne Case and Angus Deaton, "Rising mortality and morbidity in midlife", Proceedings of the National Academy of Sciences, 2015

HRS Cohort Design



Cohort changes in health (ages 51-56 in indicated year)







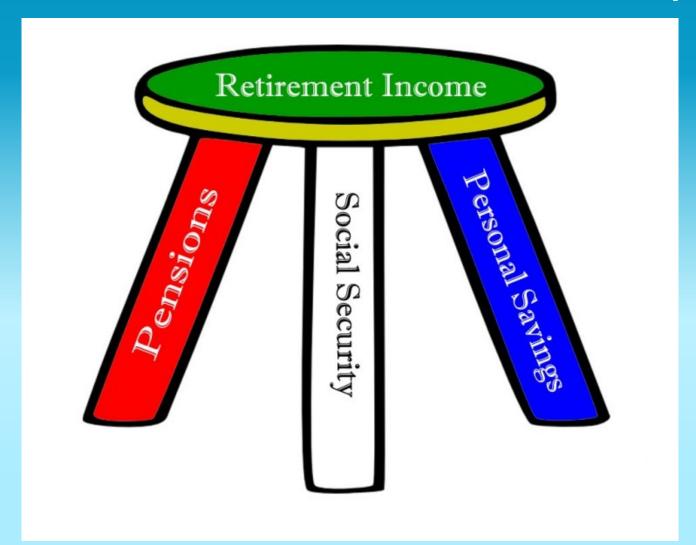
So what's happening to retirement preparation?

HRS and Retirement Security

 A key aim of HRS has always been to assess preparation for retirement of successive cohorts in their early 50s.



HRS and Retirement Security





How do we measure these legs?

- Personal savings (mostly home equity)
 - Survey self-report
- Social Security
 - Consent-based individually identified linkage to administrative records
 - + imputation for non-consenters
 - Wealth = PDV of future benefits
- Employer pensions
 - Survey + non-identified linkage to employer plan information
 - + imputation for unmatched
 - Wealth = PDV of future benefits



Pensions on Current Jobs

- For DB plans, we need to know the plan rules to estimate benefits and wealth
 - This includes terminated, frozen, and cash balance conversions
- 1992-2004 a lot of effort went into trying to obtain employer plans
 - From employers, from respondents, from the web
 - With increasingly dismal results
- For 2010, we were able to go the Department of Labor's website and obtain PDF copies of detailed 5500 filings from private employers, including attachments not in electronic versions of Form 5500



Employee Benefits Security Administration

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Filings

Form 5500/5500-SF Search

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To find actuarial information for plan year 2008, use the Pension Plan Actuarial Information Search.

You can view one plan filing at a time through this search engine based on the search criteria listed below. If you would like the raw data from all filings, including schedules, you can download Form 5500 datasets from EBSA's electronic FOIA webpage. If you are unable to find a filing that you believe has been submitted, please contact the EFAST2 Help Line at 866-463-3278.

Enter criteria in one or more of the below fields (at least one of these fields must be filled in).

Wildcard searches are automatically used in the Plan Name, Sponsor Name, and Plan Administrator fields. The wildcard is executed on the right side of the character string. For example, the search for "Widg" will show information for all plans that begin with the letters "Widg", such as "Widget Company" and "Widgard LLC". At least four letters must be entered into these fields when performing a search.

Wildcard searches are not used in fields other than Plan Name, Sponsor Name, and Plan Administrator. Only exact matches are displayed for these other fields.

For help searching for filings through this website, see the EFAST2 Filing Search Guide.

Fill in the search criteria, then click Search.



Employer Plan Match Rates, by Sector and Self-Report of Plan Type, 2004 and 2010

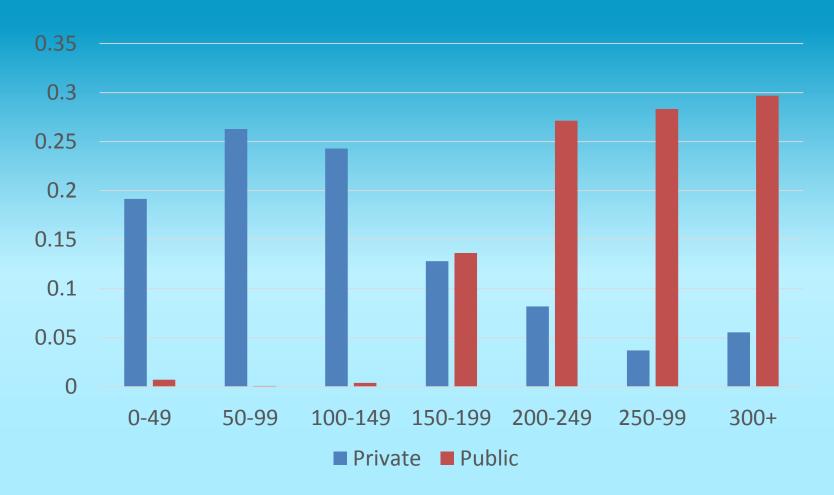
| Sector | Plan Type | 2004 | 2010 |
|---------|-----------|-------|-------|
| Private | All | 31.6% | 87.9% |
| | Any DB | 33.6% | 88.4% |
| Public | All | 88.2% | 96.9% |
| | Any DB | 92.1% | 98.3% |



Does it matter whether you match someone correctly to their own plan, or will any similar employer do?



Variation in DB plan generosity: Wealth of a simulated "typical" worker under all HRS-coded plans from 2010, by sector (\$2010 000s)





What if we relied entirely on employer match?

- As designed by Gustman and Steinmeier (and described in their book with Nahid Tabatabai on Pensions in the HRS), the HRS approach has been to rely on self-report of plan type and use employer match only when it agrees.
- With low match rates, not much alternative, even though evidence for plan type error is strong
- With high match rates, we can consider substituting employer match for self-report



HRS pension wealth in 2010, comparing self-report of plan type to employer match of type (\$000s)

| | DB wealth | | DC wealth | | <u>Total Pension</u> <u>wealth</u> | |
|------------------------|-----------|---------|-----------|---------|---------------------------------------|---------|
| Plan type agreement | Resp | Empl | Resp | Empl | Resp | Empl |
| Agree | 139,964 | 139,494 | 201,160 | 201,160 | 341,123 | 341,123 |
| Type switch | 13,698 | 14,579 | 7,085 | 20,437 | 20,783 | 35,016 |
| Drop plan | 40,468 | 28,624 | 40,128 | 22,565 | 80,596 | 51,189 |
| Add plan | 69,856 | 116,006 | 83,706 | 140,579 | 153,562 | 256,585 |
| Total | 263,986 | 298,704 | 332,079 | 384,740 | 596,065 | 683,443 |



Other evidence on incomplete selfreport

- Employer match might raise pension wealth by about 10-13% - Is that plausible?
- For persons linked to IRS-SSA-W2 records, we can compare deferred compensation on W2 with selfreport of plan participation (see Dushi and Honig, 2014)
 - Only available for linked cases, doesn't provide information on balance, doesn't capture plans not getting active contributions
- It does confirm there are DC plans missing in selfreport, and more often than not in places where employer match also finds them



Some adjustments for our cohort comparisons

- Pro-rate DB wealth based on years of service (as in Gustman, Steinmeier, and Tabatabai)
- Add in value of plans on past jobs
- Pro-rate SS wealth based on PIA formula and covered earnings up to entry into HRS
- Include IRAs and assign them to individuals
- Split other household wealth evenly in twoperson households



Components of Wealth, Cohorts aged 51-56 by Year (\$2010)

| Wealth | 1992 | 1998 | 2004 | 2010 |
|------------------|---------|---------|---------|---------|
| IRA | 15,147 | 26,409 | 30,395 | 26,238 |
| DC | 22,152 | 38,497 | 35,711 | 44,675 |
| DB | 109,856 | 108,086 | 60,549 | 35,881 |
| SSW | 104,139 | 104,154 | 120,166 | 127,313 |
| Total retirement | | | | |
| wealth | 251,294 | 277,146 | 246,821 | 234,107 |
| HH with | 176,744 | 177,530 | 217,082 | 179,699 |
| Total wealth | 428,038 | 454,676 | 463,903 | 413,806 |



The Three Legs

- Personal savings
 - Flat, but 2010 was hit by housing crisis
 - May see some growth later
- Social Security
 - Increasing with rising incomes
- Private pensions
 - Shrinking
 - Loss of DB plans not compensated by growth in DC+IRA wealth

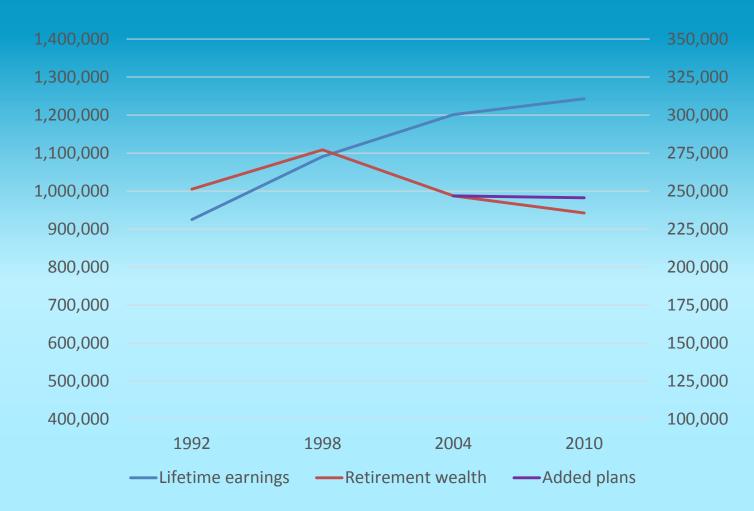


Lifetime Earnings and Retirement Wealth at Age 51-56, by Year





Lifetime Earnings and Retirement Wealth at Age 51-56, by Year





Lifetime Earnings and Retirement Wealth, by Position in Distribution of Lifetime Earnings



600,000 550,000 500,000 450,000 400,000 350,000 300,000 250,000 2010 Retirement wealth

Lower half of lifetime earnings

Upper half of lifetime earnings



Lifetime Earnings and Retirement Wealth, by Race/Ethnicity



African-American

Hispanic



Lifetime Earnings and Retirement Wealth, by Gender



Men Women



Summary and Policy Implications

The views expressed here are my own



Summary and Policy Implications-1

- Thank you, DoL, and keep up the good work!
 - Even though employer matching does not increase estimated pension wealth enough to change the aggregate path of cohort change,
 - It does make individual-level data more accurate
- Next up on the worry list: public DB plans



Summary and Policy Implications-2

- Retirement wealth has been declining, in absolute value and as a share of lifetime earnings
- The growth in DC + IRA balances has not compensated for the decline in DB values
- Can we do more to nudge DC offerings by employers and take-up by workers?



Summary and Policy Implications-3

- Social Security is a larger share of total retirement wealth for later cohorts, and becoming the only source with mandatory annuitization
- But this depends on solving solvency!
- Can someone please save Social Security?



THANK YOU!



Comments on: Cohort Changes in Social Security Benefits and Pension Wealth by Fang, Weir, and Brown

1

Irena Dushi Social Security Administration

Presented at the 18th Annual RRC Meeting, Washington DC August 4-5, 2016

The opinions expressed here are my own and do not represent the view of SSA

Key questions

2

 How financially prepared, particularly in terms of SSW and PW, are younger cohorts (MBB) of pre-retirees compared to previous cohorts?

• What are the differences in SSW and PW within cohorts such as across the income distribution?

Importance of Having SSW and PW estimates

- Provide a complete picture of the balance sheet of pre-retirees, when combined with other wealth components available in the HRS
- Allow to assess retirement preparedness of pre-retirees
- Calculating PW and SSW is time consuming and requires different assumptions. Having these wealth measures created in a consistent way from the HRS is an enormous contribution to the research community

Key takeaways

4

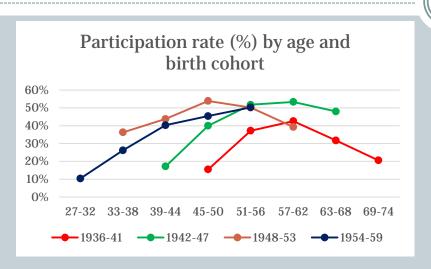
• Employer pensions (DB+DC+IRA) are providing a much lower share of total wealth for the MBB cohort than for the HRS cohort (25% vs. 34%)

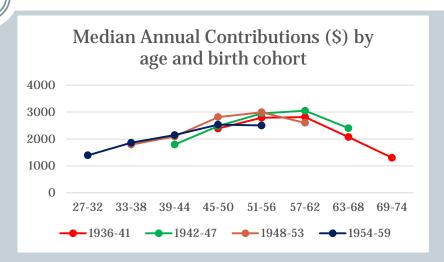
• SSW comprises a higher share of total wealth than PW for the MBB cohort (31% and 25%); the reverse is true for the HRS cohort (24% and 34%, respectively)

IRA+DC wealth much lower than DB wealth

- 5
- Stock market shocks all retirement accounts lost 31% in assets from September 07 - May 09 (Soto 2009)
- Leakages Argento, Bryant, Sabelhaus (2013) find:
 - Pre-retirement withdrawals from IRAs and DC plans are substantial, and strongly correlated with income and marital status shocks
- Job and earnings shocks Dushi and Iams (2015) using w-2 data find:
 - Even during non-recessionary periods and even when employees stay in the same job, their participation and contributions to DC plans fluctuate considerably
 - O Among workers with <u>no job change</u>: those with decreased earnings were more likely to stop or decrease contributions than those with stable earnings
 - Among workers <u>with stable earnings</u>: those who changed jobs were more likely to stop contributing than those with no job change
- → Assuming consistent and increasing contributions in order to project DC wealth at retirement is likely to overestimate DC wealth

What W-2 data tell us about DC plans

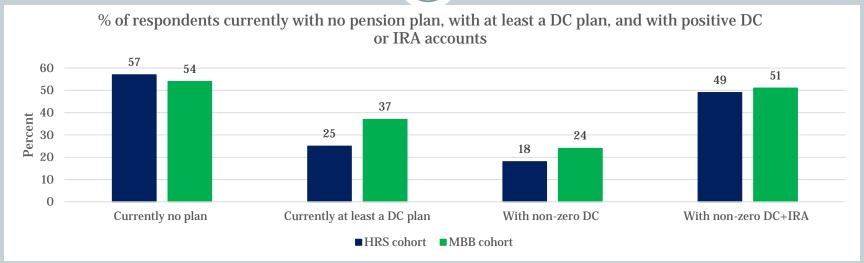


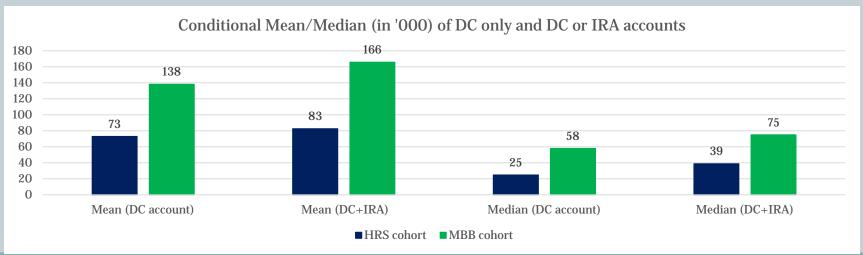


Source: Honig and Dushi (2009), estimates updated with recent data for MBB cohort

| Contributions according to W-2 record, by cohort | | | | |
|---------------------------------------------------|------------|-----------|----------|----------|
| | HRS - 1992 | WB - 1998 | EBB-2004 | MBB-2010 |
| Participation Rate % | 37 | 54 | 50 | 50 |
| # of years with contrib. in the past 10 years (%) | | | | |
| None | 69 | 57 | 47 | 48 |
| 1-5 years | 21 | 19 | 20 | 21 |
| 6-9 years | 9 | 15 | 20 | 14 |
| 10 years | 1 | 10 | 13 | 17 |
| Total \$ contributions in the past 10 years | | | | |
| Mean | 14392 | 21021 | 24936 | 26540 |
| Median | 8500 | 12437 | 13201 | 12742 |

Survey-reported DC and IRA accounts





Policy Implications



- The small amount of PW (DB+DC+IRA) among younger cohorts suggests that they will be more dependent on SS in retirement
- Younger cohorts will need to work longer and save more in voluntary DC plans in order to support their standard of living in retirement
 - Good news: Average expected probability of working full time past age 62 and 65 is higher for the MBB cohort compare to earlier cohorts
- Policy makers should continue to design policies that induce employers
 offering, employees participation and savings, such as the automatic
 plan features, as well as provisions that discourage early withdrawal

Calculating PW

- O Respondents' self-reported vs. employers' plan data since the inaccuracy of self-reported pension information may lead to inaccurate PW estimates (Gustman and Steinmeier (2004b) and Dushi and Honig (2014))
- Fang, Brown, and Weir (2016) use data from employers' Form 5500 to compare and correct self-reported plan type and calculate respective PW
 - However, the magnitude of (in)consistencies between self-reported and firm data regarding plan type is not presented in the paper
 - Also, authors do not use the W-2 data to correct for respondents misreport of DC plans
 - Are employers' reported plan data (or matching of these data) subject to errors?
 - Both respondent & employer reported DB only plan 12% had W-2 contributions
 - Employer reported DB only and employee reported Both 63% had W-2 contributions
- → W-2 records are the gold standard for DC plans, use them to determine and correct for DC plan type by looking at contributions in the current or the past few years

Calculating PW (cont.)

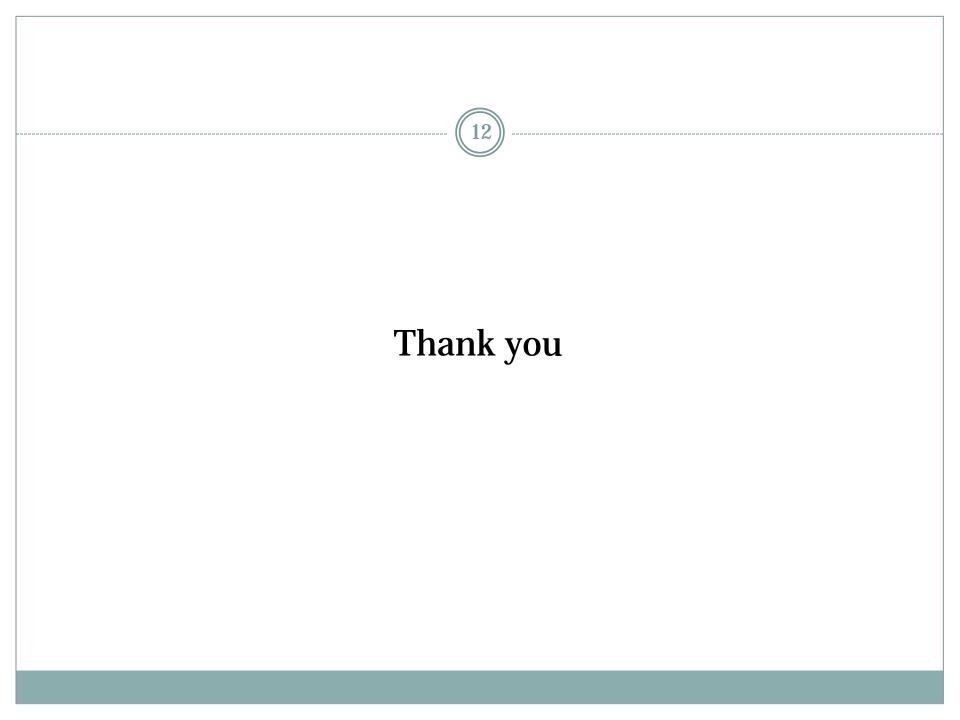


- Use survey reported earnings and employment histories to project future earnings, but when available use W-2 records to determine hire dates
 - **X** Suggestions:
 - For those with matched W-2 records use actual W-2 earnings to project future earnings, as you currently do for SSW
- Impute DC account balances based on nearest neighbor matching method
 - **×** Suggestion:
 - Use information from adjacent waves when available
 - Can also use w-2 records to validate whether and how much contributed in the past few years

Calculating SSW



- SSW estimates are not available for respondents who have claimed SS benefits
 - **×** Suggestion:
 - include the value of remaining SSW for those who have currently claimed benefits
- Calculate SSW only at the first wave when cohorts enter the survey
 - **X** Suggestion:
 - create wave-specific measures of SSW (also of PW) as information from survey reports in consecutive waves and from administrative records become available



How Does Student Debt Affect Early-Career Retirement Saving?

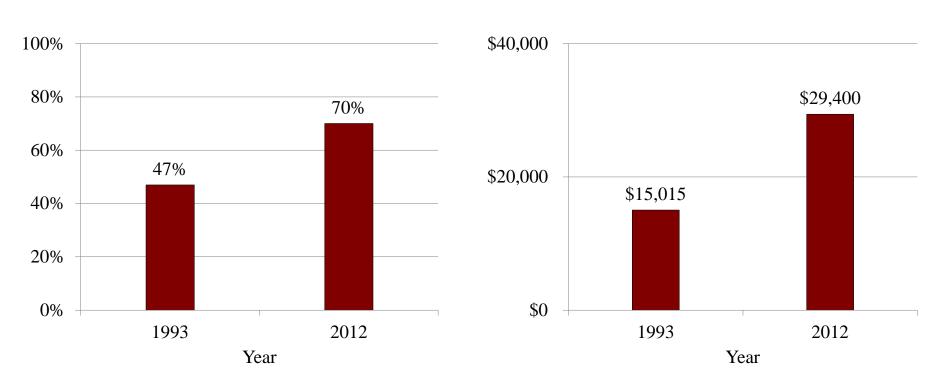
Matthew S. Rutledge
Research Economist
Center for Retirement Research at Boston College

18th Annual Meeting of the Retirement Research Consortium Washington, DC August 5, 2016



Student loan borrowing has increased greatly over the last two decades.

Proportion of Graduates with Student Loans and Average Real Loan Balance, 1993 and 2012



Note: Student loan balances in 2012 dollars.

Source: The Institute on College Access and Success. 2014. "Quick Facts about Student Debt." Oakland, CA.



How could young households handle the student debt burden?

- Cut back on consumption
- Rack up other debt
- Cut back on non-retirement savings
- Cut back on savings for retirement



Research question: How does student debt affect retirement saving?

- Joint work with Geoffrey Sanzenbacher, Francis Vitagliano.
- Not well-studied because of recency of problem.
- Most studies use the *Survey of Consumer Finances* (SCF).
 - Small sample of young households with debt.
 - Often include older households with their kids' debt.
 - Little background information on student debtors.



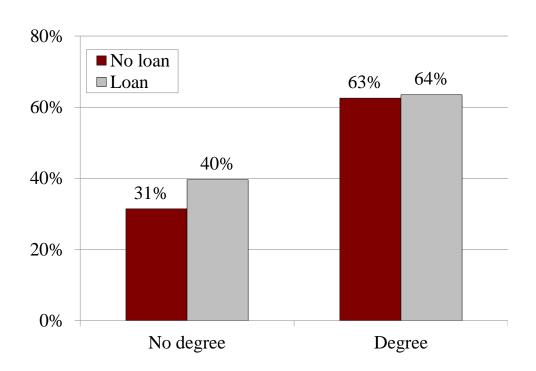
Instead, we use the data from the *National Longitudinal Survey of Youth*, 1997 Cohort.

- Key ages of observation:
 - Age 25: respondent's student loan balance
 - Age 30: household retirement assets, plan participation
- Unique controls for debtor's background
 - College quality: public, private non-profit, private forprofit
 - Parents' education and income at age 30
 - ASVAB aptitude test score



Retirement plan participation is not lower for student debtors compared to non-debtors.

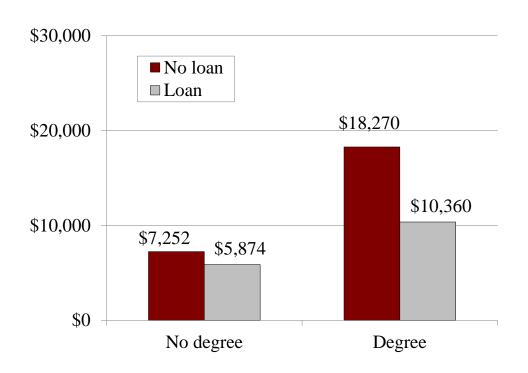
Retirement Plan Participation by Degree Status and Student Loan Status





But median assets half as large for graduates with debt relative to graduates without debt.

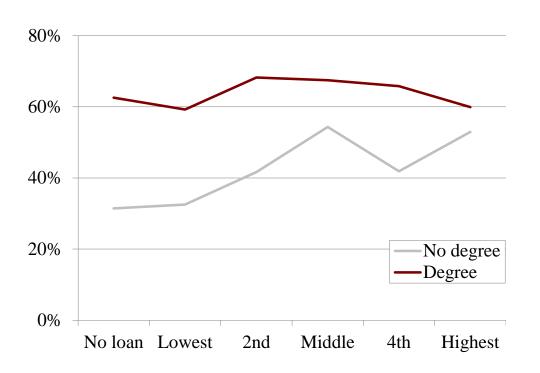
Median Retirement Assets by Degree Status and Student Loan Status





No reduction in plan participation as student loan balances increase.

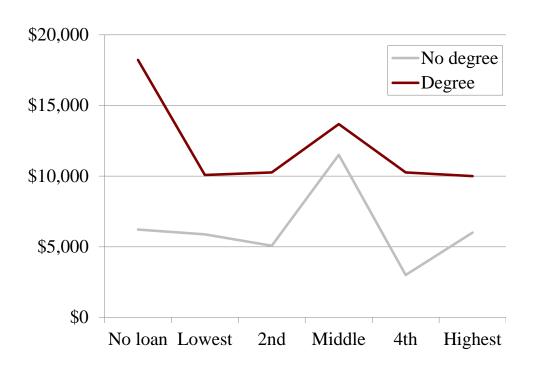
Retirement Plan Participation Rate by Degree Status and Student Loan Quintile





Median assets also do not decrease as loans increase; difference is in loans vs. no loans.

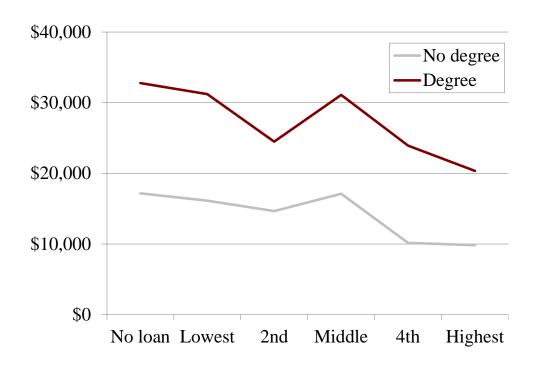
Median Retirement Assets by Degree Status and Student Loan Quintile





But mean assets do decline as loans increase.

Average Retirement Assets by Degree Status and Student Loan Quintile





Econometric model

- Outcomes of interest:
 - Participation in employer-sponsored plan
 - Take-up of an employer's pension offer
 - Retirement assets
- Key variables: indicator for having loan, ln(loan balance)
- Base specification: loan controls, earnings, firm size, demos
- Then add degree status, school quality, parental background, ASVAB percentile



No statistically significant difference in plan participation with loans or as loans grow.

Retirement Plan Participation Regression Results

| | All | | No degree | | Bachelor's | |
|----------------------------|---------|---------|-----------|---------|------------|---------|
| | (1) | (2) | (3) | (4) | (5) | (6) |
| Student loan (0/1) | 0.030 | -0.007 | 0.038 | 0.010 | -0.142 | -0.147 |
| | (0.101) | (0.103) | (0.172) | (0.176) | (0.168) | (0.173) |
| Ln real loan balance at 25 | 0.011 | 0.013 | 0.028 | 0.026 | 0.022 | 0.020 |
| | (0.011) | (0.011) | (0.021) | (0.021) | (0.017) | (0.018) |
| Sample size | 3,313 | 3,074 | 1,508 | 1,293 | 985 | 978 |

Notes: Students' variables are measured as of age 30. Regressions also include gender, marital status, presence of children, race, and Hispanic ethnicity.



Take-up of pensions also does not decrease with greater loans.

Participation Regressions among Workers Ever Offered a Plan

| | All | | No degree | | Bachelor's | |
|----------------------------|---------|---------|-----------|---------|------------|---------|
| | (1) | (2) | (3) | (4) | (5) | (6) |
| Student loan (0/1) | 0.074 | 0.001 | 0.036 | -0.002 | -0.172 | -0.171 |
| | (0.107) | (0.109) | (0.188) | (0.193) | (0.179) | (0.186) |
| Ln real loan balance at 25 | 0.026** | 0.026** | 0.033 | 0.034 | 0.029 | 0.025 |
| | (0.011) | (0.012) | (0.023) | (0.023) | (0.018) | (0.019) |
| Sample size | 2,804 | 2,626 | 1,184 | 1,025 | 881 | 876 |

Notes: Students' variables are measured as of age 30. Regressions also include gender, marital status, presence of children, race, and Hispanic ethnicity. ** p<0.05.



Assets may be lower with loans, but not statistically significantly.

Retirement Asset Regression Results

| | All | | No degree | | Bachelor's | |
|----------------------------|---------|---------|-----------|---------|------------|---------|
| | (1) | (2) | (3) | (4) | (5) | (6) |
| Student loan (0/1) | 0.298 | 0.163 | -0.244 | -0.355 | -0.381 | -0.414 |
| | (0.410) | (0.408) | (0.838) | (0.851) | (0.612) | (0.609) |
| Ln real loan balance at 25 | -0.003 | -0.009 | 0.068 | 0.055 | -0.017 | -0.009 |
| | (0.043) | (0.043) | (0.098) | (0.099) | (0.062) | (0.062) |
| Sample size | 1,472 | 1,420 | 468 | 422 | 565 | 564 |

Notes: Dependent variable is the log of retirement assets. Students' variables are measured as of age 30. Regressions also include gender, marital status, presence of children, race, and Hispanic ethnicity.



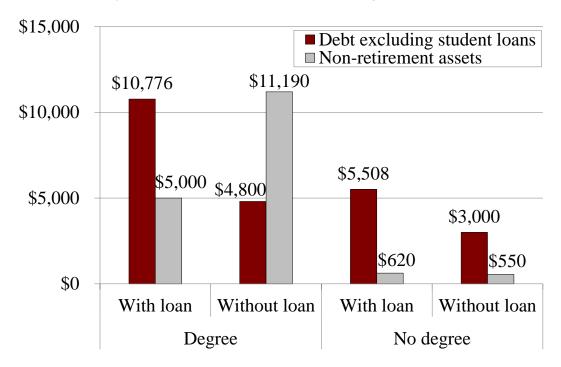
No evidence student loans affect retirement saving in alternative models, either.

- Limiting the sample to unmarried individuals.
- In asset regressions, adding number of years in a pension plan.
- Using a nonlinear function of the student loan balance.



Instead, student debtors have more non-educational debt, lower assets.

Median Non-Educational Debt and Non-Retirement Asset Levels at Age 30, by Student Loan Status and Degree Status





Conclusion

- Student loan debt clearly hurts young workers' finances.
- But student loans do not appear to hurt retirement saving by age 30.
- So disadvantage in student debt manifests in other negative financial outcomes.
- Will retirement savings still be unaffected for the coming cohorts with even larger student debt burdens?





Comments:

How Does Student Debt Affect Early-Career Retirement Saving?

Rutledge, Sanzenbacher, and Vitagliano

High debt is associated with better balance

Figure 5

Typical High-Debt Gen Xers and Millennials Tend to Have Better Balance Sheets Than Their Low-Debt Peers

High debt, income, and net worth, by generation

| Hig | h-debt |
|-----|--------|
| | Xers |

3 times higher income than low-debt peers

6.5 times higher net worth than low-debt peers

High-debt millennials

3 times higher income than low-debt peers

5.5 times higher net worth than low-debt peers

Source: Pew Survey of American Family Finances

© 2015 The Pew Charitable Trusts

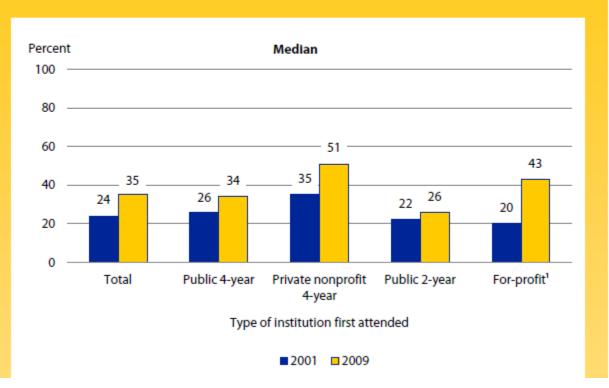
Source: The Pew Charitable Trusts, 2015, "The Complex Story of American Debt" http://www.pewtrusts.org/en/research-and-analysis/reports/2015/07/the-complex-story-of-american-debt

Non-completers

FIGURE 7.

TOTAL-FEDERAL-DEBT-TO-ANNUAL-INCOME-RATIO AMONG NONCOMPLETERS

Cumulative federal loan debt as a percentage of annual income at the 50th percentile (bars) and at the 25th and 75th percentiles (table below), by type of first institution attended: 2001 and 2009



Source: Christina Chang Wei and Laura Horn, 2013, "Federal Student Loan Debt Burden of Noncompleters" NCES http://nces.ed.gov/pubs2013/2013155.pdf

Graduate Students

FIGURE ES.1

Debt Is Higher among Graduates with Higher Degrees

Cumulative debt among degree recipients, 2012

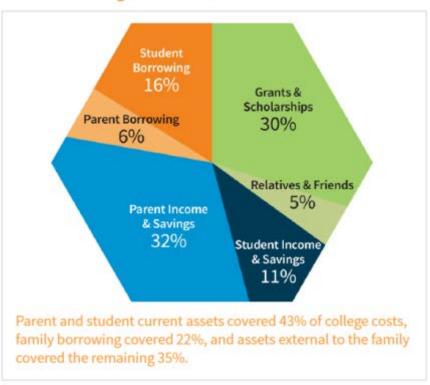


Source: Sandy Baum and Martha C. Johnson, 2015, "Student Debt: Who Borrows Most? What Lies Ahead?" Urban Institute

http://www.urban.org/research/publication/student-debt-who-borrows-most-what-lies-ahead

What about the parents?

Figure 6 — How the Typical Family Pays for College 2014-15, Average Percent of Total Cost Paid from Each Source

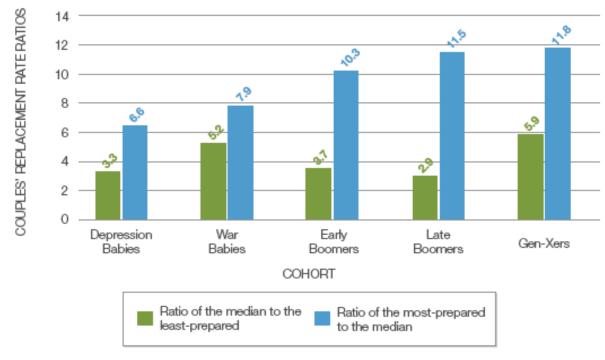


Source: Sallie Mae and Ipsos, 2015, "How America Pays for College 2015" http://news.salliemae.com/files/doc_library/file/HowAmericaPaysforCollege2015FNL.pdf

A Larger Question of Inequality

RETIREMENT READINESS HAS BECOME MORE UNEQUAL OVER TIME

FIGURE 12. REPLACEMENT RATE RATIOS BETWEEN THE MEDIAN AND THE LEAST-PREPARED, AND BETWEEN THE MOST-PREPARED AND THE MEDIAN, BY COHORT



Source: Panel Study of Income Dynamics.

Note: The replacement rates are calculated assuming retirement at age 65. The least-prepared, median, and most-prepared are represented by the replacement ratios of the 1st, 50th, and 99th percentiles, respectively.

Source: The Pew Charitable Trusts, 2013, "Retirement Security Across Generations" http://www.pewtrusts.org/~/media/legacy/uploadedfiles/pcs_assets/2013/empretirementv4051013finalforwebpdf.pdf



Conclusion

- Non-completers low loan balances
 – low human capital, employers with fewer retirement benefits
- Grad students high loan balances, but excellent prospects for repayment and retirement
- Are we worrying about the wrong age group? Is it parents we should think about?
- Larger questions around inequality



Diana Elliott Senior Research Associate Urban Institute

delliott@urban.org @dianabelliott

Marital Histories, Gender, and Financial Security in Late Mid-Life: Evidence from the Health and Retirement Study (HRS)

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What This Study is About

- How are more recent cohorts faring in terms of pre-retirement financial security?
- How have marital histories changed across cohorts?
- How do marital histories relate to measures of financial security and what does this entail for more recent cohorts approaching retirement?
 - Is there variation by gender?
- We use multiple regression models using 4 cohorts from the Health and Retirement Study.

Marital History and Financial Security in Later Life across Cohorts

- Continuous marriage = \$
 - Resource pooling, specialization, economies of scale (Becker 1973, 1981)
 - Encourage saving behavior
 - Economic advantages accrue over marital duration (Addo and Lichter, 2013)
 - Social Security policy (Meyer and Herd, 2007)

- "Marital history" incorporates not only current marital status but also the timing, sequencing, and duration of marital statuses across an individual's life course.
 - This project: past and present marital statuses.

Thinking about Marital Histories and Gender Together

- The financial benefits of marriage differ by gender
 - Men: marriage premiums and fatherhood premiums for wages for men (Glauber 2008; Killewald 2012; Korenman and Neumark 1991)
 - Women: wage penalties for both marriage and motherhood (Budig and England 2001)
- Marital dissolution likewise varies in its economic consequences for men and women
 - Women are more economically disadvantaged following divorce or widowhood (Lavelle and Smock 2012; Meyer and Herd 2007)
 - Men are more likely to remarry (Shafer and James 2013)
- Combined with other gender differences in economic and demographic factors, differences in the economic consequences of marriage increase women's financial vulnerability in later life (Favreault 2005; Fisher et al. 2009)

Marriage is Changing

- In more recent birth cohorts, there have been increases in:
 - Divorce
 - Cohabitation—pre-marriage or as a substitute
 - Never-marriage
 - Children born outside of marriage, multi-partnered fertility
 - "Marriage-go-round"—de-partnering and re-partnering (Cherlin, 2010)
- Marriage has also changed in additional ways related to:
 - Selectivity
 - Division of labor by gender within and outside of the household
 - Expansion to same-sex couples
- What do these demographic shifts in marriage mean for the financial security of men and women approaching retirement?

Data: Health and Retirement Study

- We examine four birth cohorts at ages 51-56.
- 1992: Sample of Americans born 1931-1941 (aged 51-61) ("Original HRS Cohort")
 - We focus on a subset born 1936-1941 (aged 51-56 in 1992)
- 1998: War Babies born 1942-1947 added
- 2004: Early Boomers born 1948-1953 added
- 2010: **Mid-Boomers** born 1954-1959 added
- Period effects? Cohort effects?

Elaborated Marital History Categories (n=17)

Partnered

- Continuously married
- Currently remarried
 - Divorced once
 - Divorced twice
 - Divorced three or more times
 - Widowed once
- Currently cohabiting
 - Never married
 - Divorced or widowed once
 - Divorced or widowed twice

Single

- Never married
- Separated
 - From first marriage
 - From second marriage
 - From third marriage
- Divorced
 - Once
 - Twice
 - Three or more times
- Widowed once
- Divorced and widowed at least once

Selected Marital History Categories (%) by Cohort, Ages 51-56 (Weighted)

| | Original HRS (b. 1936-1941) (Wave 1, 1992) | War Babies (b. 1942-1947) (Wave 4, 1998) | Early Boomers (b. 1948-1953) (Wave 7, 2004) | Middle Boomers (b. 1954-1959) (Wave 10, 2010) |
|---------------------------------------|--------------------------------------------------|------------------------------------------------|---------------------------------------------------|-----------------------------------------------------|
| Continuously married | 52.4% | 49.2% | 45.3% | 43.7% |
| Currently remarried, divorced once | 14.4% | 15.7% | 16.5% | 15.7% |
| Currently remarried, divorced twice | 2.8% | 3.6% | 3.9% | 4.2% |
| Currently, remarried, divorced thrice | 0.9% | 0.8% | 1.1% | 0.8% |

Selected Marital History Categories (%) by Cohort, Ages 51-56 (Weighted)

| | Original HRS (b. 1936-1941) (Wave 1, 1992) | War Babies (b. 1942-1947) (Wave 4, 1998) | Early Boomers (b. 1948-1953) (Wave 7, 2004) | Middle Boomers (b. 1954-1959) (Wave 10, 2010) |
|---------------------------|--------------------------------------------------|------------------------------------------------|---------------------------------------------------|-----------------------------------------------------|
| Never married | 3.8% | 4.3% | 5.4% | 7.6% |
| Single, divorced once | 8.6% | 10.6% | 9.1% | 9.1% |
| Single, divorced twice | 2.7% | 3.5% | 4.5% | 4.0% |
| Single, divorced thrice/+ | 1.0% | 0.8% | 1.1% | 1.0% |

Financial Security Measures (Mean or %) by Cohort, Ages 51-56 (Weighted)

| | Original HRS (b. 1936-1941) (Wave 1, 1992) | War Babies (b. 1942-1947) (Wave 4, 1998) | Early Boomers (b. 1948-1953) (Wave 7, 2004) | Middle Boomers (b. 1954-1959) (Wave 10, 2010) | | |
|------------------------------------------------|-------------------------------------------------------|------------------------------------------------|---------------------------------------------------|-----------------------------------------------------|--|--|
| Net Wealth (includes | Net Wealth (includes primary and secondary residence) | | | | | |
| Negative | 4.6% | 4.6% | 6.2% | 10.9% | | |
| Zero | 2.9% | 2.0% | 2.2% | 3.2% | | |
| Positive | 92.5% | 93.3% | 91.6% | 85.9% | | |
| Positive wealth, 2016\$, top-coded | \$366,029 | \$432,249 | \$485,511 | \$354,743 | | |
| Earnings, full-time workers, 2016\$, top-coded | \$50,972 | \$52,211 | \$60,868 | \$55,640 | | |

| Multinomial Logistic Regression | Negative v. Positive | Wealth; RRR (SE) | Zero v. Positive W | ealth; RRR (SE) |
|------------------------------------|----------------------|------------------|--------------------|-----------------|
| Original HRS (ref.) | | | | |
| War Babies | 0.89 | (0.10) | 0.75 + | (0.10) |
| Early Baby Boomers | 1.17 | (0.12) | 0.91 | (0.10) |
| Middle Baby Boomers | 2.43 *** | (0.21) | 1.52*** | (0.20) |
| Continuously married (ref.) | | | | |
| Remarried, divorced once | 1.44 *** | (0.15) | 0.93 | (0.20) |
| Remarried, divorced twice | 1.67** | (0.29) | 0.83 | (0.40) |
| Remarried, divorced thrice | 2.40** | (0.73) | 0.00 | (0.00) |
| Cohabiting, never married | 1.43 | (0.45) | 7.02*** | (1.90) |
| Cohabiting, divorced/widowed once | 1.90** | (0.35) | 1.96 | (0.60) |
| Cohabiting, divorced/widowed twice | 1.30 | (0.46) | 1.35 | (0.80) |
| Single, never married | 3.30*** | (0.40) | 6.96*** | (1.10) |
| Single, divorced once | 2.70*** | (0.28) | 4.56*** | (0.70) |
| Single, divorced twice | 2.95** | (0.44) | 4.31*** | (1.00) |
| Single, divorced thrice/+ | 2.23* | (0.62) | 4.23*** | (1.70) |

Notes: +p<0.10, *p<0.05, **p<0.01, ***p<0.001. Covariates in model but not shown here include gender, age, race, education, number of children ever-born, labor force status, total household income, self-rated health, and count of chronic conditions. N=15,296

| OLS Positive Wealth (logged, 99% top-coded, 2016\$) | Unstandardized beta | (SE) |
|-----------------------------------------------------|---------------------|--------|
| Original HRS (ref.) | | |
| War Babies | 0.06+ | (0.04) |
| Early Baby Boomers | 0.05 | (0.04) |
| Middle Baby Boomers | -0.22*** | (0.04) |
| Continuously married (ref.) | | |
| Remarried, divorced once | -0.16*** | (0.04) |
| Remarried, divorced twice | -0.40*** | (0.07) |
| Remarried, divorced thrice | -0.42** | (0.15) |
| Cohabiting, never married | -0.87*** | (0.15) |
| Cohabiting, divorced/widowed once | -0.48* | (0.09) |
| Cohabiting, divorced/widowed twice | -0.43** | (0.15) |
| Single, never married | -1.00*** | (0.07) |
| Single, divorced once | -0.84*** | (0.05) |
| Single, divorced twice | -1.22*** | (0.08) |
| Single, divorced thrice/+ | -1.67*** | (0.14) |

Notes: +p<0.10, *p<0.05, **p<0.01, ***p<0.001. Covariates in model but not shown here include gender, age, race, education, number of children ever-born, labor force status, total household income, self-rated health, and count of chronic conditions. N=13,442.

| OLS FT Earnings (logged, 99% top-coded, 2016\$) | Men, unstandardized beta (SE) | | Women, unstandardized beta (SE) | |
|-------------------------------------------------|-------------------------------|--------|---------------------------------|--------|
| Original HRS (ref.) | | | | |
| War Babies | 0.05 | (0.03) | 0.00 | (0.03) |
| Early Baby Boomers | 0.03 | (0.03) | 0.07* | (0.03) |
| Middle Baby Boomers | 0.07* | (0.03) | 0.16*** | (0.03) |
| Continuously married (ref.) | | | | |
| Remarried, divorced once | -0.02 | (0.03) | 0.03 | (0.04) |
| Remarried, divorced twice | -0.03 | (0.06) | 0.02 | (0.07) |
| Remarried, divorced thrice | -0.16 | (0.12) | 0.10 | (0.15) |
| Cohabiting, never married | -0.60*** | (0.14) | -0.20 | (0.14) |
| Cohabiting, divorced/widowed once | -0.10 | (0.09) | -0.10 | (0.08) |
| Cohabiting, divorced/widowed twice | -0.22 | (0.14) | -0.10 | (0.14) |
| Single, never married | -0.35*** | (0.06) | -0.10 | (0.05) |
| Single, divorced once | -0.12* | (0.05) | 0.04 | (0.04) |
| Single, divorced twice | -0.01 | (0.08) | 0.10 | (0.06) |
| Single, divorced thrice/+ | -0.09 | (0.14) | 0.06 | (0.11) |

Notes: +p<0.10, *p<0.05, **p<0.01, ***p<0.001. Covariates in model but not shown here include age, race, education, number of children ever-born, labor force status, hours worked per week, self-rated health, and count of chronic conditions. N (men)=4,300; N (women)=3,949.

Summary

- At ages 51-56, Middle Baby Boomers (b. 1954-1959) differ from those in the original HRS Cohort (b. 1936-1941) in several ways.
 - MBBs less likely to be continuously married, more likely to be never-married, cohabit, and have complex marital histories.
 - MBBs are more likely to have negative or zero wealth
 - MBBs who have positive wealth have lower levels of wealth
 - MBBs (and EBBs) working full-time have higher real earnings.
- Marital history is more strongly related to wealth than earnings.
- Gender differences by cohort for earnings but not wealth.
- Past marital history, not just present marital status, matter especially with regards to wealth.
- Period effect of Great Recession, cohort effect of 1970s stagflation

Implications of Increasing Marital Complexity

- Policy
 - Social Security
 - Medicare and Medicaid
- Individuals and Families
 - Financial behaviors and decision-making
 - Family caregiving
- Increasing heterogeneity, increasing inequality in later life? (gender, race, class)

Discussion

MARITAL HISTORIES, GENDER, AND FINANCIAL SECURITY IN LATE MID-LIFE

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Highlights

- IMPORTANT TOPIC
 - How does the decline in marriage affect preparedness for retirement?
- HRS reveals new insights
 - Detailed data on wealth
 - Detailed data on marital histories
 - Multiple cohorts nearing retirement

Highlights

- KEY RESULTS
 - Big change in family structure across HRS cohorts
 - Decline in marriage, increase in divorce
 - both matter
 - trends will be magnified further for future cohorts
 - along with decline in fertility
 - Non-long-term marriage is associated with lower wealth, at ages 51-56 Though not lower earnings for women
 - Most recent HRS cohort has lower wealth, at ages 51-56

Extensions

- EMPIRICAL STRATEGY
- Can be viewed as differences-in-differences
 - Older cohorts serve as control for younger cohorts
 - Differences in marriage within cohorts identify wealth, earnings effects
 - Do older cohorts help control for underlying trends
 - e.g. rising inequality, changes in fertility
 - Do older cohorts help control for contemporaneous shocks
 - e.g. housing crisis, Great Recession
 - Did marriage patterns change because of other factors also affecting wealth
 - Tests for pre-trends, placebo outcomes

Extensions

- POLICY IMPLICATIONS
 - Marriage law
 - State laws affect likelihood, cost of divorce, division of marital property
 - Old-age programs
 - How to treat marriage, divorce in Social Security, Medicaid, SSI